

# Arthroscopic Latarjet yields better union and prevention of instability compared to arthroscopic bony Bankart repair in shoulders with recurrent anterior instability: a systematic review

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Title: Arthroscopic Latarjet yields better union and prevention of instability compared to
 arthroscopic bankart repairs in shoulders with recurrent anterior instability: a systematic
 review

#### 4 Abstract

**Purpose:** To determine whether arthroscopic Latarjet procedure arthroscopic Latarjet or
arthroscopic bony Bankart repair arthroscopic bony Bankart repair provide better outcomes in
terms of rates of recurrent instability, non-union and complications, as well as clinical scores
and range of motion.

9 Methods: An electronic literature search was performed using PubMed, Embase®, and
10 Cochrane databases, applying the following keywords: "Arthroscopic bony bankart" OR
11 "Arthroscopic osseous bankart" AND "Arthroscopic Latarjet" OR "Arthroscopic coracoid
12 Bone block".

Results: The systematic search returned 1,465 records, of which 29 were included 13 14 (arthroscopic bony Bankart repair, n=16; arthroscopic Latarjet, n=13). 37 datasets were 15 included for data extraction, on 1483 shoulders. Compared to arthroscopic Latarjet, 16 arthroscopic bony Bankart repair had significantly higher instability rates (0.14; CI, 0.10-17 0.18; vs 0.04; CI, 0.02–0.06), significantly lower union rates (0.63; CI, 0.28–0.91 vs 0.98; CI, 18 0.93-1.00), and significantly lower pain on VAS (0.42; CI, 0.17-0.67 vs 1.17; CI, 0.96-1.38). There were no significant differences in preoperative glenoid bone loss, follow-up, 19 complication rate, ROWE score, ASES score, external rotation, and anterior forward 20 elevation between arthroscopic Latarjet and arthroscopic bony Bankart repair. 21

Conclusion: Compared to arthroscopic Latarjet, arthroscopic bony Bankart repair results in
significantly (i) higher rates of recurrent instability (14% vs 4%), (ii) lower union rates (63%
vs 98%), but (iii) slightly lower pain on VAS (0.45 vs 1.17). There were no differences in
complication rates, clinical scores, or postoperative ranges of motion.

26 Level of Evidence: Level IV

27 Keywords: Arthroscopic Latarjet procedure, bony Bankart procedure, anterior instability,
28 outcomes

#### 29 Introduction

30 Anterior shoulder instability often requires surgical treatment, particularly in young athletes who wish to resume sports [3,46,50]. The most common surgical treatments involve either 31 32 labral or capsular repair in shoulders with no glenoid defects [34,51], or bone-block procedures in shoulders with glenoid defects [13,26,35]. The Latarjet technique is among the 33 34 most popular bone-block procedures, as it proved effective at preventing recurrent dislocations [1,2,12,21,25,27,39] whether performed as an open or an arthroscopic procedure 35 [15,33]. The Latarjet technique is however technically demanding and therefore has a 36 37 considerable learning curve and requires alteration of native scapular anatomy [5,9,15].

38

39 More recently, arthroscopic bony Bankart repair arthroscopic bony Bankart repair[49] was 40 introduced as a less invasive alternative to arthroscopic Latarjet, to treat anterior shoulder 41 instability with traumatic glenoid defects, where the bone fragment is still present and could 42 be sutured back onto the glenoid [4]. While arthroscopic bony Bankart repair demonstrated 43 satisfactory bone union rates, it is associated with higher rates of recurrent instability compared to arthroscopic Latarjet [6,39,52,54], and may be insufficient to treat shoulders 44 with irreparable ligamentous damage, attritional glenoid bone loss, or major bone 45 deficiencies [15]. 46

47

There is no consensus on the optimal treatment for recurrent anterior shoulder instability with traumatic glenoid defects. Therefore, the purpose of this study is to determine whether arthroscopic Latarjet or arthroscopic bony Bankart repair provide better outcomes for this population in terms of rates of recurrent instability, non-union and complications, as well as clinical scores and range of motion.

#### 53 Materials and Methods

The protocol for this systematic review was submitted to PROSPERO prior to commencement (registration number BLINDED) and conforms to the principles outlined in the handbook of the Cochrane Collaboration [19], along with the guidelines established by the Preferred Reporting Items for Systematic Reviews and Meta-analysis (PRISMA) [37].

58

#### 59 *Search strategy*

The authors conducted a structured electronic literature search on 06 July 2022 using the 60 61 PubMed, Embase®, and Cochrane Central Register of Controlled Trials databases, applying the following keywords: "Arthroscopic bony bankart" OR "Arthroscopic osseous bankart" 62 63 AND "Arthroscopic Latarjet" OR "Arthroscopic coracoid Bone block". The full search 64 strategy is presented in Appendix I. The search was limited to articles published between 2002 and 2022. After removal of duplicate records, each of two researchers (BLINDED) 65 screened the titles and abstracts to determine the suitability for the review using the following 66 67 predefined eligibility criteria:

68

## 69 Inclusion criteria

Comparative or non-comparative studies on patients with recurrent anterior shoulder
 instability that underwent either primary arthroscopic Latarjet or primary arthroscopic
 bony Bankart repair with or without adjuvant procedures

- Studies that measured preoperative glenoid bone loss using computed tomography
(CT) scans

Studies that report pre- and post-operative clinical outcomes in terms of, recurrence of
 instability, non-union, complication rates, range of motion, or clinical scores

77

#### 78 *Exclusion criteria:*

Narrative or systematic reviews, case reports, expert opinions, editorials, letters to
 editors, computer simulations, cadaver or laboratory studies

Studies published in languages other than English, French, German, Italian, Spanish,
and Dutch to avoid translation errors.

83

#### 84 *Study selection*

Studies that met the eligibility criteria during title and abstract screening underwent full-text screening by two researchers (BLINDED) and any disagreement was first discussed between the researchers, and if required, a third researcher (BLINDED) resolved any disagreement. The reference lists of studies for full text review were searched, and an expert (BLINDED) was consulted to identify further relevant studies that may not have been captured by the database searches.

91

### 92 Data extraction and quality assessment

Data extraction was performed by two researchers (BLINDED) independently and their 93 94 results were compared to ensure accuracy. Where there was disagreement in the documented value, the true value was ascertained by simultaneous review of the data in question by both 95 96 researchers. The following information was extracted from the included studies; author(s), 97 journal, year of publication, level of evidence, country where study was performed, conflicts of interest and funding declaration. Furthermore, the following clinical data was extracted: 98 type of surgery, subgroup size, sex, age, preoperative glenoid bone loss, union rates (instead 99 100 of non-union rates which were inconsistently defined and reported), recurrent instability rates (defined as recurrence of subluxation and/or dislocation), post-operative complication rates, 101

range of motion (ROM; ER1, arm at 0° of abduction; ER2, arm at 90° of abduction; AFE,
anterior forward elevation), clinical scores, and pain on VAS scale.

104

105 Methodological quality of the eligible studies was assessed by two researchers (BLINDED) 106 according to the Joanna Briggs Institute (JBI) checklist, to appraise the reporting quality (10 107 items). Where there was disagreement between the researchers, consensus was achieved by 108 discussion and review.

109

#### 110 Statistical analysis

When available in the original articles, outcomes were tabulated: continuous outcomes were 111 112 reported as means, standard deviations, and ranges, while categorical outcomes were reported as proportions. Pooled estimates were calculated if mean and standard deviations were 113 available for least 3 studies reported. When the standard deviation was not reported, the range 114 was used to calculate standard deviation using a method outlined by Hozo et al.[22]. 115 Heterogeneity was evaluated by visual inspection of forest plots and using the I<sup>2</sup> statistic and 116 its connected  $\chi^2$  test, to provide a measure of the degree of inconsistency across studies. 117 Pooled estimates of raw means and their 95% confidence interval (CI) were calculated using 118 119 a random-effects model framework. Pooled estimates of proportions and their 95% confidence intervals (CI) were calculated via Freeman-Tukey double arcsine transformation 120 121 using inverse-variance weighting within a random-effects model framework. P-values <0.05122 were considered statistically significant. Statistical analyses were performed using R version 4.1.3 (R Foundation for Statistical Computing, Vienna, Austria) using the meta package. 123

#### 124 **Results**

125 The systematic search returned 1,465 records, of which 563 were duplicates, leaving 902 for title and abstract screening (Figure 1). A total of 786 studies were excluded by examining 126 127 their titles and/or abstracts because: 500 were not relevant, 223 were editorials/conference proceedings etc., 29 were reviews, 17 were published before 2002, and 17 were published in 128 languages other than those defined in the exclusion criteria. A further 87 studies were 129 excluded after full-text review. This left 29 eligible studies, [1,7,8,11,14,16-18,23,24,29-130 32,38,40-45,47,48,53,55-59] of which 16 assessed arthroscopic bony Bankart repair, 13 131 assessed arthroscopic Latarjet (Table 1 and 2). All studies reported preoperative glenoid bone 132 loss: 21 reported mean preoperative glenoid bone loss, of which 4 created subgroups of 133 134 glenoid bone loss and reported the mean loss of each subgroup, while 8 used a threshold of 135 glenoid bone loss as a part of their inclusion criteria. Of the 29 eligible studies, 37 datasets 136 were included for data extraction, with a total of 1,483 shoulders (arthroscopic bony Bankart 137 repair, n=874; AL, n=609) (Table 3 and 4). Of the 37 included datasets, 12 datasets reported 138 on concomitant Hill-Sachs (arthroscopic bony Bankart repair, n=7 and arthroscopic Latarjet, 139 n=5) lesions.

140

Of the 29 eligible studies, 2 were prospective and 27 were retrospective. Of the 29 studies, 3 did not report criteria for inclusion and 2 did not specify the presenting sites demographic information. Furthermore, it was unclear in 4 studies whether the study had complete inclusion of participants. Overall, all included studies had good to excellent scored according to the JBI checklist (Table 5).

147

148 Preoperative glenoid bone loss

<sup>141</sup> *Quality assessment* 

Of the 37 included datasets, preoperative glenoid bone loss was reported as part of their inclusion criteria in 7 and was reported as a mean in 26 datasets (arthroscopic bony Bankart repair, n=19; arthroscopic Latarjet, n=7). There was no significant difference in mean preoperative glenoid bone loss between shoulders treated by arthroscopic bony Bankart repair (17.6%; CI, 13.7–21.6; I<sup>2</sup>=99%) or arthroscopic Latarjet (18.9%; CI, 13.5–24.3; I<sup>2</sup>=98%) (Figure 2).

155

156 Follow-up

Of the 37 included datasets, mean follow-up was reported in 24 (arthroscopic bony Bankart
repair, n=13; arthroscopic Latarjet, n=11). There was no significant difference in mean
follow-up between shoulders treated by arthroscopic bony Bankart repair (49.3 months; CI,
27.9–70.8; I<sup>2</sup>=100%) or arthroscopic Latarjet (29 months; CI, 20.7–37.3; I<sup>2</sup>=99%).

161

#### 162 *Recurrent instability*

Of the 37 included datasets, rates of recurrent instability were reported in 19 (arthroscopic
bony Bankart repair, n=13; arthroscopic Latarjet, n=6). There was a statistically significant
higher rate of instability following arthroscopic bony Bankart repair (0.14; CI, 0.10–0.18;
I<sup>2</sup>=49%) compared to arthroscopic Latarjet (0.04; CI, 0.02–0.06; I<sup>2</sup>=0%) (Figure 3).

167

#### 168 Union

169 Of the 37 included datasets, union was reported in 14 (arthroscopic bony Bankart repair, n=5; 170 arthroscopic Latarjet, n=9). All studies assessed union using postoperative CT scans, and 171 there was a statistically significant lower union rate following arthroscopic bony Bankart 172 repair (0.63; CI, 0.28–0.91; I<sup>2</sup>=96%) compared to arthroscopic Latarjet (0.98; CI, 0.93–1.00; 173 I<sup>2</sup>=70%) (Figure 4).

## 175 *Postoperative complications*

176 Of the 37 included datasets, complication rates were reported in 15 (arthroscopic bony 177 Bankart repair, n=3; arthroscopic Latarjet, n=12), of which 3 reported rates instead of 178 absolute values [23,57]. There was no statistical difference in complication rate following 179 arthroscopic bony Bankart repair (0.01; CI, 0.00–0.07;  $I^2=0\%$ ) compared to arthroscopic 180 Latarjet (0.03; CI, 0.00–0.06;  $I^2=47\%$ ) (Figure 5).

181

#### 182 *Clinical scores*

183 There were no significant differences between arthroscopic bony Bankart repair and
184 arthroscopic Latarjet, in terms of ROWE score (92.78 vs 92.47), and ASES score (94.26 vs
185 92.16). However, arthroscopic bony Bankart repair demonstrated significantly lower pain on
186 VAS (0.42) compared to arthroscopic Latarjet (1.17; p<0.01).</li>

187

## 188 *Range of motion*

- 189 There were no significant differences between arthroscopic bony Bankart repair and
- arthroscopic Latarjet, in terms of ER1 ( $59.5^{\circ}$  vs  $57.26^{\circ}$ ), and AFE ( $167.93^{\circ}$  vs  $170.90^{\circ}$ ).

#### 191 Discussion

192 The most important findings of this meta-analysis are that, compared to arthroscopic Latariet, 193 arthroscopic bony Bankart repair results in significantly (i) higher rates of recurrent 194 instability (14% vs 4%), (ii) lower union rates (63% vs 98%), but (iii) slightly lower pain on VAS (0.45 vs 1.17). There were however no significant differences in complication rates, 195 196 clinical scores, or postoperative ranges of motion. The null hypothesis that the two procedures would result in comparable outcomes is therefore refuted. The clinical relevance 197 198 of these findings is that, while arthroscopic bony Bankart repair demonstrated comparable 199 clinical scores to arthroscopic Latarjet, the latter might be more suitable in patients prone to 200 recurrent instability or non-union. It is worth noting, however, that the present meta-analysis 201 did not account for the surgeon experience and learning curves, which should also be 202 considered when choosing the surgical technique.

203

204 The present systematic review included studies that had good to excellent scores according to 205 the JBI checklist. A previous systematic review by Longo et al.[36] compared outcomes of all surgical treatments for traumatic anterior glenohumeral instability with glenoid and/or 206 207 humeral bony defects. Longo et al.[36] found that shoulders with glenoid bone loss <25% 208 were most frequently treated with arthroscopic Bankart repair (50%) or arthroscopic bony Bankart repair (32%), while shoulders with glenoid bone loss >25% were most frequently 209 210 managed by open reconstruction with bone graft (37%), arthroscopic Bristow procedure 211 (26%), or open Latarjet procedure (OLP) (22%). The present meta-analysis found no significant difference in mean preoperative glenoid bone loss between shoulders treated with 212 213 arthroscopic bony Bankart repair (17.6%; CI, 13.7–21.6) versus arthroscopic Latarjet (18.9%; CI, 13.5–24.3). However, it remains unclear whether the choice of procedure should be based 214 on the extent of glenoid bone loss, or rather on the size of the remaining bone fragment 215

216 present that could be sutured back onto the glenoid. It is worth noting that arthroscopic bony 217 Bankart repair may be insufficient to treat shoulders with irreparable ligamentous damage, 218 attritional glenoid bone loss, or major bone deficiencies [15], where bone block procedures 219 may be the only treatment option. A recent systematic review by Hurley et al.[28] compared outcomes of ABR alone versus ABR with remplissage versus OLP, to treat anterior shoulder 220 221 instability in patients with concomitant Hill-Sachs lesions. Hurley et al.[28] found higher rates of recurrent instability following ABR alone (16.8%) compared to OLP (7.0%) and 222 223 ABR with remplissage (3.2%). A study by Horinek et al. [20] compared OLP with ABR with 224 remplissage, for anterior shoulder instability and found no difference in terms of clinical 225 scores, return to sport, and satisfaction. In the present meta-analysis, only one study 226 performed arthroscopic bony Bankart repair with remplissage (in 11 of 312 patients) [47] but 227 did not report the outcomes of this subgroup. Furthermore, out of the 29 included studies in 228 the present meta-analysis, Hill-Sachs lesions were reported in 15 studies, and their prevalence 229 ranged from 24% to 100%. It was not possible to perform a subgroup analysis of the 230 outcomes of this pathology as the type of Hill-Sachs lesion was not consistently reported, 231 which could affect the rate of recurrent instability [10,36].

232

233 The results of the present meta-analysis should be interpreted with the following limitations in mind. All included studies were case series, although the level of quality was good to fair. 234 235 Furthermore, the thresholds of glenoid bone loss reported were heterogenous across studies 236 which makes it difficult to determine whether the choice for arthroscopic bony Bankart repair 237 or arthroscopic Latarjet should be based on the extent of preoperative glenoid bone loss. This 238 meta-analysis did not evaluate Hill-Sachs lesions which could have affected recurrent instability rates [10,36] as the specific lesions were not consistently reported. Additionally, 239 there is need for a more standardised method of reporting outcomes as a great number of 240

different clinical scores were reported, making it difficult to quantitively compare studies.
Finally, it is worth noting that, while there was no statistically significant difference in
follow-up between the two techniques, the mean follow-up for arthroscopic bony Bankart
repair was almost 2 years longer than for arthroscopic Latarjet, which could introduce bias in
the comparisons.

246

#### 247 *Conclusion*

This systematic review and meta-analysis found that, compared to arthroscopic Latarjet, arthroscopic bony Bankart repair results in significantly (i) higher rates of recurrent instability (14% vs 4%), (ii) lower union rates (63% vs 98%), but (iii) slightly lower pain on VAS (0.45 vs 1.17). There were however no significant differences in complication rates, clinical scores, or postoperative ranges of motion.

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- 456 Figure legend
- **Figure 1:** Flowchart of study inclusion
- **Figure 2:** Forest plot on glenoid defect size
- **Figure 4:** Forest plot on recurrent instability
- **Figure 5:** Forest plot on union
- **Figure 6:** Forest plot on postoperative complications
- 463 Supplementary material:
- **Figure 1:** Forest plot on follow-up
- **Figure 2:** Forest plot on ROWE score
- **Figure 3:** Forest plot on ASES score
- **Figure 4:** Forest plot on pain on VAS
- Figure 5: Forest plot on external rotation 1
- **Figure 6:** Forest plot on active forward elevation
- 470 Appendix 1: Search strategy

# 472 F1



473

474 F2

Author, Year	n Shoulders	Mean	SD	Glen	oid defect size	MRAW	95%-CI
Surgery = Bony Bankart							
Di Giacomo, 2022 (≥13.5%)	40	16.10	2.7000		+	16.10	[15.26; 16.94]
Di Giacomo, 2022 (<13.5%)	23	23.90	2.4000		+	23.90	[22.92; 24.88]
Nakagawa, 2022 (a) (13.5-20%	b) 60	16.20	2.1000		+	16.20	[15.67; 16.73]
Nakagawa, 2022 (a) (>20%)	36	23.80	3.2000		+	23.80	[22.75; 24.85]
Guo, 2021	25	27.50	9.3000		+	→ 27.50	[23.85; 31.15]
Godin, 2019	13	22.50	8.6711			22.50	[17.79; 27.21]
Nakagawa, 2019	69	9.50	7.2500	-	<b>—</b>	9.50	[ 7.79; 11.21]
Park, 2018 (<10% )	30	5.30	4.8000	-+-		5.30	[ 3.58; 7.02]
Park, 2018 (10-15%)	12	16.40	4.7000			16.40	[13.74; 19.06]
Park, 2018 (15-20%)	26	19.20	6.1000			19.20	[16.86; 21.54]
Park, 2018 (>20%)	17	27.00	5.1000			- 27.00	[24.58; 29.42]
Nakagawa, 2017 (a)	93	11.30			t	11.30	
Nakagawa, 2017 (b)	113	9.50		1		9.50	
Kitayama 2015	34	20.40			ĩ	20.40	
Nakagawa, 2015	81	16.80			1	16.80	
Kim, 2014 (<12.5%)	16	8.80	2.0000	+	a	8.80	[7.82; 9.78]
Kim, 2014 (12.5%-25%)	18	17.50	2.7500		-+-	17.50	[16.23; 18.77]
Millett, 2013	12	29.00	9.3095			+> 29.00	[23.73; 34.27]
Park, 2012	31	14.10	8.6000		_ <b>—</b>	14.10	[11.07; 17.13]
Random effects model	749				$\bigcirc$	18.31	[14.35; 22.28]
Heterogeneity: $I^2 = 99\%$ , $\tau^2 = 49$ .	1806, <i>p</i> < 0.01						
<u> </u>							
Surgery = Latarjet							
Hurley, 2022	26	19.10	4.7000			19.10	[17.29; 20.91]
Shao, 2022	30	21.60	8.6000			21.60	[18.52: 24.68]
Bonnevialle, 2021	17	10.60	4.5000	1)	<del></del>	10.60	[ 8.46; 12.74]
Calvo, 2021	51	18.80			1	18.80	
Hurley, 2021	30	18.60	4,7000			18.60	[16.92: 20.28]
Ali. 2020	33	15.00	5.0000			15.00	[13.29: 16.71]
Zhu, 2017 (a)	46	23 60	2 7000		+	23.60	[22 82: 24 38]
Zhu, 2017 (b)	52	23.30	2.8000		+	23.30	[22.54: 24.06]
Random effects model	285				$\sim$	18.86	[14.51: 23.21]
Heterogeneity: $l^2 = 97\%$ , $\tau^2 = 21$	3161. p < 0.01						[
Lice gonory in or ro, the Line			1	1			
			(	) 5 1	0 15 20 25	30	

**F3** 

Author, Year	Recurrence	Cohort Size		Proportion	95%-CI
Surgery = Bony Bankart					
Di Giacomo, 2022 (≥13.5%)	3	40		0.07	[0.02; 0.20]
Di Giacomo, 2022 (<13.5%)	3	23	• • •	0.13	[0.03: 0.34]
Nakagawa, 2022 (a) (13.5-20%)	7	60		0.12	[0.05; 0.23]
Nakagawa, 2022 (b)	18	80	<b>i</b> >	0.23	[0.14; 0.33]
Nakagawa, 2019	15	69		0.22	[0.13: 0.33]
Shah, 2018 (>25%)	2	22		0.09	[0.01; 0.29]
Nakagawa, 2017 (a)	22	93		0.24	[0.15; 0.34]
Nakagawa, 2017 (b)	23	113		0.20	[0.13; 0.29]
Kitayama 2015	1	34	•	0.03	[0.00; 0.15]
Nakagawa, 2015	12	81		0.15	[0.08; 0.24]
Kany, 2013	1	23		0.04	[0.00; 0.22]
Millett, 2013	1	12	· · · · · · · · · · · · · · · · · · ·	0.08	[0.00; 0.38]
Park, 2012	2	31		0.06	[0.01; 0.21]
Random effects model		681	$\sim$	0.14	[0.10; 0.18]
Heterogeneity: $I^2 = 49\%$ , $\tau^2 = 0.005$	3, <i>p</i> = 0.02				
0					
Surgery = Latarjet	0	00		0.00	10 04 0 051
Hurley, 2022	2	26		0.08	[0.01; 0.25]
Calvo, 2021	3	51		0.06	[0.01; 0.16]
Hurley, 2021	2	30		0.07	[0.01; 0.22]
Ali, 2020	1	33		0.03	[0.00; 0.16]
Bolleau, 2019 (>20%)	4	121		0.03	[0.01; 0.08]
Kordasiewicz, 2019 (>10%)	3	90		0.03	[0.01; 0.09]
Random effects model		351	$\diamond$	0.04	[0.02; 0.06]
Heterogeneity: $I^- = 0\%$ , $\tau^* = 0$ , $p = 0$	0.81	Ĩ			

0 0.05 0.1 0.15 0.2 0.25 0.3

**F**4

Author, Year	Union	Cohort Size			Proportion	95%-CI
Surgery = Bony Bankart						
Nakagawa, 2022 (b)	48	80	a		0.60	[0.48; 0.71]
Guo, 2021	25	25			1.00	[0.86; 1.00]
Nakagawa, 2019	17	69			0.25	[0.15; 0.36]
Nakagawa, 2017 (a)	26	93			0.28	[0.19; 0.38]
Park, 2012	26	31		-	0.84	[0.66; 0.95]
Random effects model		298			0.63	[0.28; 0.91]
Heterogeneity: $I^2 = 96\%$ , $\tau^2 = 0$	.1527, p	< 0.01				•
Surgery = Latarjet						
Shao, 2022	29	30		+	0.97	[0.83; 1.00]
Bonnevialle, 2021	12	17		_	0.71	[0.44; 0.90]
Calvo, 2021	51	51			1.00	[0.93; 1.00]
Wang, 2020 (screw) (>20%)	6	6			1.00	[0.54; 1.00]
Wang, 2020 (button) (>20%)	6	6			1.00	[0.54; 1.00]
Boileau, 2019 (>20%)	115	121			0.95	[0.90; 0.98]
Kordasiewicz, 2019 (>10%)	81	90			0.90	[0.82; 0.95]
Zhu, 2017 (a)	46	46			1.00	[0.92; 1.00]
Zhu, 2017 (b)	52	52			1.00	[0.93; 1.00]
Random effects model		419		$\diamond$	0.98	[0.93; 1.00]
Heterogeneity: $I^2 = 70\%$ , $\tau^2 = 0$	.0168, p	< 0.01				-
		0	0.2 0.4 0.6	0.8 1		

482 F5

Author, Year	Complications	Cohort Size	Proportion	95%-CI
Surgery = Bony Bankart				
Guo. 2021	0	25	0.00	[0.00: 0.14]
Kitayama 2015	2	34	0.06	[0.01: 0.20]
Kim, 2014 (<12,5%)	0	16 -	→ 0.00	[0.00: 0.21]
Random effects model		75 <	0.01	[0.00: 0.07]
Heterogeneity: $I^2 = 0\%$ , $\tau^2 = 0.0$	0015, <i>p</i> = 0.37			[]
Surgery = Latarjet		1 01274022 and		
Hurley, 2022	2	26 *	→ 0.08	[0.01; 0.25]
Shao, 2022	0	30	0.00	[0.00; 0.12]
Calvo, 2021	0	51 ⊫	0.00	[0.00; 0.07]
Hurley, 2021	2	30	→ 0.07	[0.01; 0.22]
Zeng, 2021 (>15%)	4	37 *	→ 0.11	[0.03; 0.25]
Ali, 2020	5	33+	→ 0.15	[0.05; 0.32]
Wang, 2020 (screw) (>20%)	0	6	→ 0.00	[0.00; 0.46]
Wang, 2020 (button) (>20%)	0	6	→ 0.00	[0.00; 0.46]
Boileau, 2019 (>20%)	5	121	0.04	[0.01: 0.09]
Zhu, 2017 (a)	1	46	0.02	[0.00: 0.12]
Zhu, 2017 (b)	0	52 -	0.00	[0.00: 0.07]
Dumont, 2014 (>25%)	3	64	0.05	[0.01: 0.13]
Random effects model		502	0.03	[0.00: 0.06]
Heterogeneity: $l^2 = 47\%$ , $\tau^2 = 0$	0.0065, p = 0.04			
and a second				
		0 0.05 0.1 0.1	5 0.2	

**T1** 

# Table 1: Study characteristics of the studies on arthroscopic Latarjet

Year		Author	Journals	Declare COIs	Have COI	Declare funding	Have funding	Level of evidence	Country
Lat	arjet								
	2022	Hurley	Arthrosc Sports Med Rehabil	Y	Y	Y	Y	III	USA
	2022	Shao	Arthroscopy	Y	Ν	Y	Y	IV	China
	2021	Bonnevialle	AJSM	Y	Y	Ν	NR	III	France
	2021	Calvo	AJSM	Y	Y	Ν	NR	IV	Spain
	2021	Hurley	Arthrosc Sports Med Rehabil	Y	Y	Ν	NR		USA
	2021	Zeng	BMC Muscoskeletal Disorders	Y	Ν	Y	Y		China
	2020	Ali	Arthroscopy	Y	Ν	Ν	NR	III	Turkey
	2020	Wang	Orthop. Surgery	Y	Ν	Y	Ν		China
	2019	Boileau	Arthroscopy	Y	Y	Ν	NR	IV	France
	2019	Kordasiewicz	KSSTA	Y	Ν	Y	Ν		Poland
20	17 (a)	Zhu	AJSM	Y	Ν	Y	Y	IV	China
20	17 (b)	Zhu	Arthroscopy	Y	Y	Y	Ν	III	China
	2014	Dumont	AJSM	Y	Y	Ν	NR	IV	France

**T2** 

Year	Author	Journals	Declare COIs	Have COI	Declare funding	Have funding	Level of evidence	Country
Bankart								
2022	Di Giacomo	AJSM	Y	Υ	Y	Y	III	Italy and USA
2022 (a)	Nakagawa	AJSM	Y	Ν	Ν	NR	IV	Japan
2022 (b)	Nakagawa	Arthroscopy	Y	Ν	Ν	NR	III	Japan
2021	Guo	OJSM	Ν	NR	Ν	NR	IV	China
2019	Godin	AJSM	Y	Y	Y	Y	IV	USA
2019	Nakagawa	OJSM	Y	Ν	Ν	NR	III	Japan
2018	Park	AJSM	Y	Y	Y	Y	III	Korea
2018	Shah	Shoulder and Elbow	Y	Ν	Y	Ν		UK
2017 (a)	Nakagawa	AJSM	Y	Ν	Ν	NR	III	Japan
2017 (b)	Nakagawa	OJSM	Y	Ν	Ν	NR	III	Japan
2015	Kitayama	JBJS-Am	Y	Ν	Y	Ν	IV	Japan
2015	Nakagawa	AJSM	Y	Ν	Ν	Ν	III	Japan
2014	Kim	AJSM	Y	Y	Ν	NR	IV	Korea
2013	Kany	Tech Should Surg	Y	Ν	Ν	NR		India/France
2013	Millett	AJSM	Y	Y	Y	Y	IV	USA
2012	Park	Arthroscopy	Y	Ν	Ν	NR	IV	South Korea

Table 2: Study characteristics of studies on arthroscopic bony Bankart repair

# **T3**

	Studies			Patient demographics									
Year	Author	Subgroups	Number of shoulders	Men (%)	Women (%)	Age (mean) ±SD	Mean follow-up (months) ±SD	(Range)	BMI (mean) ±SD	Smoking (%)	Dominant (%)	Hill-Sachs leisons (%)	
Latarje	t												
2022	Hurley		26	81%	19%	$32\ \pm 13$	52	(24 – 90)				100%	
2022	Shao		30	70%	30%	$27 \pm 6$	$38 \pm 3$				80%	100%	
2021	Bonnevialle		17	82%	18%	$22 \pm 5$				23%	47%	59%	
2021	Calvo		51	94%	6%	$30 \pm 8$					55%		
2021	Hurley		30	83%	17%	$32 \pm 12$	46					100%	
2021	Zeng		37	100%	0%	$25 \pm 5$	13	(8 – 22)					
2020	Ali		33	88%	12%	$30 \pm 7$	30	(24 – 50)			55%	24%	
2020	Wang	screw	6	100%	0%	$23 \pm 6$	25	(10 – 53)	$23.4 \pm 2$				
		button	6	100%	0%	$23 \pm 5$			$24.9\ \pm 5$				
2019	Boileau		121	79%	21%	27	26	(24 – 27)		31%	65%		
2019	Kordasiewicz		90	89%	11%	$26 \pm 6$	$24 \pm 7$	(13 – 50)			61%		
2017(a)	Zhu		46	78%	22%	$32 \pm 10$					83%		
2017 (b)	Zhu		52	81%	19%	$32 \pm 10$	28	(24 – 42)			69%		
2014	Dumont		64	86%	14%	29	76						

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Table 4: Patient demographics of studies on arthroscopic bony Bankart repair Studies Patient demographics Mean follow-up (months) Hill-Sachs leisons (%) Number of shoulders Dominant (%) Women (%) Age (mean) Subgroups Men (%) (Range) (Range) Author (ear ₽SD TSD Bankart 2022 Di Giacomo ≥13.5% 40 100% 0% 26 (19 - 35) 60 (58 - 68)<13.5% 23 100% 0% 26 (19 - 35) 60 2022 (a) Nakagawa Overall 89% 96 11% 29 13.5% to  $\geq 20\%$ 60 85% 15%  $24\ \pm 10$ >20% 36 94% 6%  $22\ \pm 7$ 2022 (b) Nakagawa 80 29 ±11 (24 - 67) 16% Overall <13.5% 38  $18 \pm 3$ (14 - 26)≥13.5% 42 19 ±4 (14 - 29) 2021 Guo (24 - 60) 72% 28% 43 ±15 41 25 2019 Godin 13 100% 0% 40 81 (61 - 108) 39% 100% 2019 Nakagawa 69 100% 0% 18 (24 (14 - 27)15% 85 2018 Park\* Overall 91% 9%  $26\ \pm 10$ 71% <10% 30 10-15% 12 15-20% 26 >20% 17 2018 Shah 95% (>25%) 22 5% 30 (17 - 66) 28 (24 - 35)45% 2017 (a) Nakagawa 93 100% 0% 17 (14 - 27) 62% 2017 (b) Nakagawa 113 89% 11% 18 100% 2015 Kitayama 34 (15 - 34)91% 9% 24 74 2015 Nakagawa 81 90% 10% 23 (14 - 69) (17 - 41)2014 Kim <12.5% 100% 0% 25 75% 16 34 12.5%-25% 18 89% 11% 32 (19 - 43)35 61% 2013 Kany 26 88% 12% 27 (15 - 39) 74% 2013 Millett 15 87% 13% 44 (24 - 70)32 (24 - 53)47% 90% (13 - 51) 2012 Park 31 94% 6% 23 (15 - 55) 31 \*11 of 223 shoulders has additional remplissage due to engaged Hill-Sachs lesions

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	-	-	-	-	-	-	-	-	-	-
Park et al. (2012)	≻	$\mathbf{x}$	$\mathbf{x}$	Υ	Y	$\succ$	$\succ$	$\mathbf{x}$	$\mathbf{x}$	$\mathbf{x}$
Millet et al. (2013)	~	Υ	Υ	Y	D	×	×	٢	λ	Υ
Kany et al. (2013)	~	$\succ$	$\sim$	Υ	×	$\succ$	$\succ$	$\succ$	$\succ$	X
Kim et al. (2014)	~	$\succ$	$\mathbf{x}$	γ	$\mathbf{x}$	$\stackrel{\scriptstyle \star}{}$	$\stackrel{\scriptstyle \star}{}$	$\succ$	$\succ$	Υ
(2102) . In 15 awagabab	z	$\succ$		Y	Y	$\stackrel{\scriptstyle \scriptstyle \times}{}$	$\succ$	$\succ$	$\succ$	Y
Kituyama et al. (2015)	~	Y	Υ	Υ	Y	×	$\succ$	$\scriptstyle{\succ}$	Y	Y
Vakagawa et al. (2017b)	z	Y		Y	Y	$\succ$	$\succ$	$\succ$	$\stackrel{\scriptstyle \star}{}$	$\stackrel{\scriptstyle \star}{}$
Nakagawa et al. (2017a)	z	$\scriptstyle{\succ}$		Y	¥	$\succ$	$\succ$	$\succ$	$\scriptstyle{\succ}$	$\succ$
Shah et al. (2018)	~	$\mathbf{x}$	Y			$\succ$	$\succ$	$\stackrel{\scriptstyle \scriptstyle \times}{}$	$\supset$	$\stackrel{\scriptstyle \star}{}$
Park et al. (2018)	~	Υ	٢	Y	Y	$\stackrel{\scriptstyle \star}{}$	¥	Y	٢	Υ
Vakagawa et al. (2019)	~	$\sim$	٢	Y	Y	×	×	$\mathbf{x}$	$\sim$	Y
(500 ct al. (2019)	~	$\succ$	٢	Y	Y	$\mathbf{x}$	$\mathbf{x}$	$\sim$	$\succ$	Y
Guo et al. (2021)	~	$\mathbf{x}$	Y	Y	Y	$\stackrel{\scriptstyle \scriptstyle \times}{}$	$\succ$	$\sim$	$\succ$	Y
Vakagawa et al. (2022b)	~	$\mathbf{x}$	۲	Y	۲	$\succ$	$\succ$	$\succ$	$\succ$	۲
Nakagawa et al. (2022a)	~	$\sim$	Y	Y	$\stackrel{\scriptstyle \star}{}$	$\succ$	$\succ$	$\succ$	$\succ$	$\succ$
Di Giacomo et al. (2022)	~	$\sim$	$\stackrel{\scriptstyle \scriptstyle \times}{}$	Y	$\stackrel{\scriptstyle \scriptstyle \times}{}$	$\succ$	$\succ$	$\succ$	$\succ$	$\succ$
Dumont et al. (2014)	~	$\mathbf{x}$	$\stackrel{\scriptstyle \scriptstyle \star}{}$	¥	$\stackrel{\scriptstyle \star}{}$	$\succ$	$\succ$	$\stackrel{\scriptstyle \scriptstyle \times}{}$	$\stackrel{\scriptstyle \scriptstyle \times}{}$	$\mathbf{x}$
Zhu et al. (2017b)	~	$\sim$	$\sim$	$\mathbf{x}$	$\mathbf{x}$	$\succ$	$\succ$	$\succ$	$\sim$	$\mathbf{x}$
Zhu et al. (2017a)	~	$\sim$	$\mathbf{x}$	$\sim$	$\mathbf{x}$	$\succ$	$\succ$	$\succ$	$\stackrel{\scriptstyle \scriptstyle \times}{}$	$\sim$
Koradsiewicz et al. (2019)	~	Y	Y	Y	Y	$\succ$	$\succ$	$\sim$	Y	$\mathbf{x}$
Boileau et al. (2019	~	Υ	Υ	Υ	Y	$\mathbf{x}$	$\succ$	$\succ$	Y	Y
Wang et al. (2020)	~	Y	Y	Y		$\succ$	$\succ$	$\succ$	z	Y
Ali et al. (2020)	~	Y	Y	Y	Y	$\succ$	$\succ$	$\stackrel{\scriptstyle \scriptstyle \times}{}$	${\sim}$	$\mathbf{x}$
Zeng et al. (2021)	7	٢	٢	Y		×	$\stackrel{\scriptstyle \scriptstyle \times}{}$	٢	z	۲